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The Risk of Flawed Inference in Evolutionary Studies When Detectability Is Less than One

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ABSTRACT: Addressing evolutionary questions in the wild remains a challenge. It is best done by monitoring organisms from birth to death, which is very difficult in part because individuals may or may

not be resighted or recaptured. Although the issue of uncertain detection has long been acknowledged in ecology and conservation biology, in evolutionary studies of wild populations it is often assumed that detectability is perfect. We argue that this assumption may lead to flawed inference. We demonstrate that the form of natural selection acting on body mass of sociable weavers is altered and that the rate of senescence of roe deer is underestimated when not accounting for a value of detectability that is less than one. Because mark-recapture models provide an explicit way to integrate and reliably model the detection process, we strongly recommend their use to address questions in evolutionary biology.

Keywords: capture-recapture, Cormack-Jolly-Seber, fitness function, individual covariates, selection gradients, senescence.

Evolutionary biologists are often interested in addressing questions in natural populations. The increasing availability of longitudinal data collected during long-term studies in the wild makes it possible to understand evolutionary processes better than ever. In particular, evolutionary biologists are showing growing interest in performing empirical tests of evolutionary theories through the study of ongoing evolution in natural settings (Hoekstra et al. 2001; Kingsolver et al. 2001). Although studies of evolution in the wild are necessary to highlight processes that occur in the natural environment that cannot be accurately mimicked in laboratory conditions, they also contain inherent methodological problems compared to laboratory experiments, some of which are too often neglected. In particular, measuring fitness traits in the field is difficult. Estimating lifetime fitness would ideally require continual monitoring of each individual in the population from birth to death. In practice, however, individuals may or may not be seen (or recaptured) at various times in their lifetime, thereby raising the issue of a detectability value that is less than one (e.g., Lebreton et al. 1992). For example, when the last encounter with an individual occurs before the end of the study, its date of death will be unknown but will be estimated as the date the individual was last observed. Although this problem has long been acknowledged in ecology and conservation biology (Williams et al. 2002), it has surprisingly been overlooked by

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evolutionary biologists (Clobert 1995). Most often, detection is assumed to be certain so that standard methods such as generalized linear models or survival analysis can be conveniently used (Skalski et al. 2005).

Two objectives that have recently received a growing level of attention but that suffer from this problem of nonperfect detectability are (i) demonstrating microevolution within a population where selection on an evolving trait is measured (Hoekstra et al. 2001; Kingsolver et al. 2001) and (ii) investigating survival senescence and its evolution in the wild (Bronikowski and Promislow 2005; Williams et al. 2006). Here, we use two long-term data sets to illustrate the risk of flawed inference in these two types of evolutionary studies when detectability is less than unity. In both examples, we proceed in two steps. First, we conduct a naive analysis considering all individuals as being detected with certainty. Second, we adopt a mark-recapture (MR) modeling approach that explicitly accounts for the detection process (Lebreton et al. 1992); therefore, demographic parameter estimates are unbiased. Although the MR approach has been advocated to address evolutionary questions, this statistical method remains underused (Clobert 1995; Cooch et al. 2002; Brown and Brown 2004). By comparing the results of the naive and MR analyses, we show that the selection on body mass of birds is incorrectly inferred and that the rate of senescence of roe deer is underestimated.

Material and Methods

Natural Selection on Body Mass of Sociable Weavers

From 1993 to 2000, sociable weavers (*Philetairus socius*) were captured with mist nets and individually banded with numbered metal bands at Benfontein Game Farm, Northern Cape Province, South Africa. A total of 435 birds of known age that had first been ringed and weighed as juveniles were used in this analysis. Covas et al. (2002) provide further details on the MR protocol. We investigated the relationship between body mass and survival of weavers in two ways after accounting for possible age and year variation. First, using a naive approach, we assumed that detection probability (the probability of capturing an individual alive and present in the study area) was 1. We calculated selection gradients following the approach of Lande and Arnold (1983). Selection gradients measure the action of selection on a trait. More precisely, in a selection experiment, they measure the difference in the mean phenotype (i.e., mean body mass) in the selected group and the mean phenotype in the entire population before selection, holding all other traits statistically constant. Selection gradients can be estimated using the slope coefficients of a model where fitness (or a fitness component)

is a function of one or several phenotypic traits. Such relationships are also called fitness functions and are not necessarily linear: they describe the strength and the form of natural selection (stabilizing, directional, or disruptive). We calculated time to death (or equivalently, survival, a proxy of fitness in that study) as the occasion following that when an individual was last captured. Although several studies have used linear regression (e.g., Björklund and Señar 2001), logistic regression is more appropriate with binary dependent variables such as survival (Janzen and Stern 1998) and has been widely used (e.g., Husak 2006; Calsbeek and Smith 2007). Using logistic regression, we first analyzed the data using the *R* program (Ihaka and Gentleman 1996) in what we call the naive approach. We then followed recommendations by Kingsolver and Smith (1995) and used MR models to estimate the relationship between survival and body mass while accounting for a detection probability <1 . Program MARK (White and Burnham 1999) was used to estimate the likelihood of the possible relationships between survival and body mass and to estimate survival and recapture parameters. On the basis of recent analyses of the data (Covas et al. 2002; Gimenez et al. 2006), we tested for age and time dependence in survival and detection probabilities. In both the naive and the MR analyses, two models were fitted that specified linear and linear-quadratic shapes of the relationship linking survival to body mass. We discriminated between directional selection (i.e., a monotonic increase of survival with increasing body mass) and stabilizing selection (i.e., a peak of survival at some intermediate body mass) by comparing models involving a linear effect versus those that involve both linear and quadratic effects (Lande and Arnold 1983; Phillips and Arnold 1989), using the Akaike Information Criterion (AIC; Burnham and Anderson 2002): the lower the AIC value, the better the model is supported by the data. We additionally calculated "AIC weights" to quantify the likelihood that a model is the best relative to a predefined set of alternative models (Burnham and Anderson 2002). In both analyses, body mass measurements were standardized (Gimenez et al. 2006).

Senescence in Survival of Roe Deer

From 1978 to 2005, roe deer (*Capreolus capreolus*) were captured with drive nets in a 26-km² enclosure at Chizé, France. A total of 432 deer (214 females and 218 males) were used in this analysis. The age of each deer was known because they were all caught as fawns. Gaillard et al. (1993) provide further details on the MR protocol. We investigated actuarial senescence patterns (i.e., the age-related increase in mortality rate; Williams et al. 2006) in deer by estimating age-specific variation in survival. Actuarial senescence patterns in roe deer are well described using a

Gompertz curve (Gaillard et al. 2004), which assumes that the mortality hazard $\mu(x)$ is exponentially related to age x as $\mu(x) = A \exp(bx)$, where A and b are usually considered to be intercept and slope, respectively (e.g., Service 2004). We thus considered this model to be our reference. As in the previous example, a naive approach was adopted first, assuming that the probability of recapturing a deer at any stage of its reproductive life was 1. We estimated the age at death as the number of years between the first and last capture (e.g., Moya-Laraño 2002; Morbey et al. 2005). We analyzed these data using the *R* program (Ihaka and Gentleman 1996) using parametric survival modeling (Klein and Moeschberger 2003), a classical approach in senescence studies (Pletcher 1999). Individuals that were still alive at the end of the study were censored, as were some marked individuals that were removed from the population because they were injured during captures (~3%; Gaillard et al. 1993) or because they were translocated to other populations as a part of the management of this enclosed population (although translocated roe deer [~80–100 per year] were preferentially unmarked animals; see Gaillard et al. 1993 for further details). We then used an MR modeling approach to estimate age-dependent survival probabilities. On the basis of previous studies (e.g., Gaillard et al. 2004) we tested for time dependence in both survival and capture probabilities. We used program MARK (White and Burnham 1999) to estimate the parameters of the Gompertz model (Gaillard et al. 2004) and the recapture probabilities. In both analyses, males and females were analyzed separately (Gaillard et al. 1993).

Results

Natural Selection on Body Mass of Sociable Weavers

The best-fitting model included a survival probability constant over age and time, whereas detection probabilities varied over time. The linear model was selected by the naive analysis, suggesting that directional selection acts on body mass of weavers (table 1). Survival was predicted to be maximal (>0.25) for body mass <23 g and decreased in heavier birds (fig. 1, *dashed line*). When the assumption of perfect detectability was relaxed, the MR analysis revealed that estimated recapture probability values were clearly <1 and also were highly variable with time (0.12–

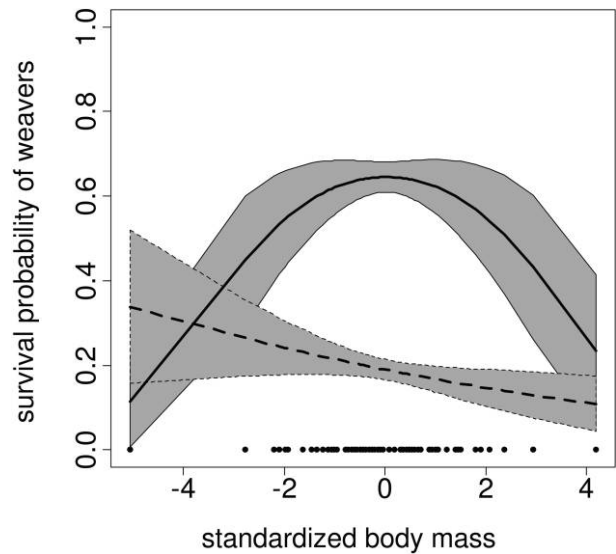


Figure 1: Relationship between survival and body mass in sociable weavers obtained by a mark-recapture analysis (*solid line*) and a naive analysis assuming perfect detection (*dashed line*). Filled circles represent body mass values, and shaded areas represent 95% confidence intervals.

0.83; see appendix table A1 for annual estimates of detection probabilities). In contrast to the naive approach, the MR approach favored the quadratic model, indicating that body mass was under stabilizing selection (table 1). Survival was maximal around the mean phenotype (fig. 1, *solid line*), that is, a body mass between 26 and 31 g leading to survival value >0.60. Overall, stabilizing selection was more than twice as plausible as directional selection (AIC weights ratio, 0.709 : 0.291), whereas the naive analysis incorrectly selected the reverse pattern, that is, suggesting directional selection was about twice as plausible as stabilizing selection (AIC weights ratio, 0.664 : 0.336). This discrepancy between the naive and the MR analyses suggests that the curvilinear relationship between survival and body mass may be masked either by a direct correlation between body mass and detection probability (e.g., Janzen et al. 2000; but additional analyses did not provide such evidence) or by an indirect relation of both parameters with a third life-history trait, such as age or body size. The observed pattern of stabilizing selection probably results

Table 1: Model selection for the relationship between survival and body mass of sociable weavers: naive approach assuming perfect detectability versus mark-recapture (MR) approach

Fitness function	Naive analysis			MR analysis		
	AIC	AIC weight	No. parameters	AIC	AIC weight	No. parameters
Linear + quadratic	427.857	.336	4	927.336	.709	10
Linear	426.495	.664	3	929.109	.291	9

Note: AIC, Akaike Information Criterion.

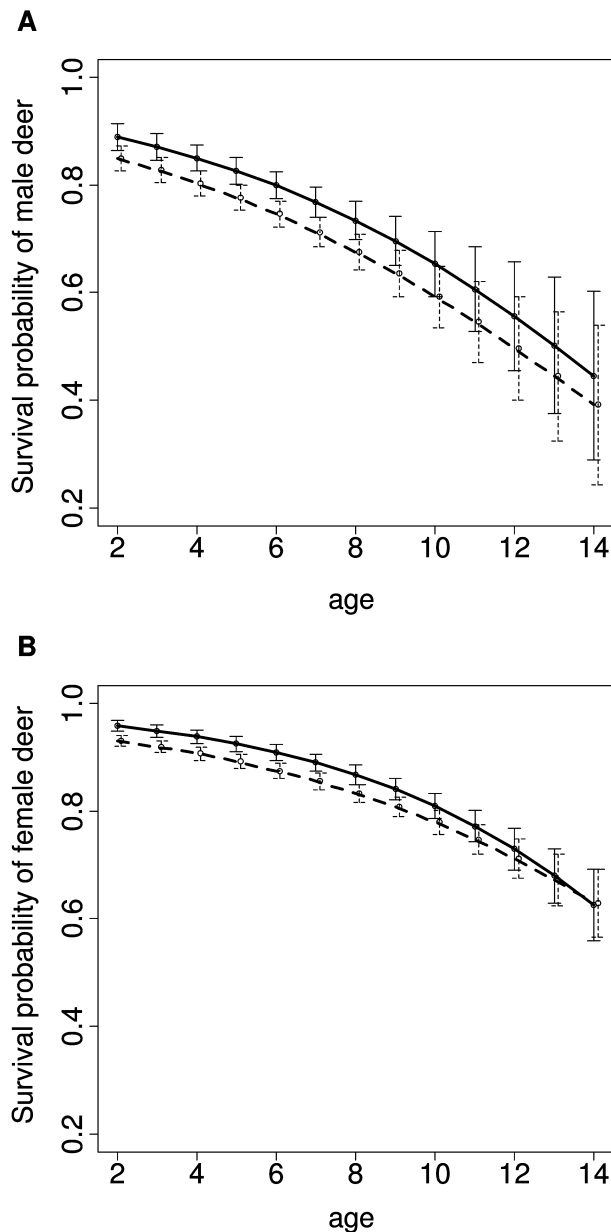


Figure 2: Relationship between survival and age in (A) male and (B) female roe deer obtained by a mark-recapture analysis (solid line) and a naive analysis assuming perfect detection (dashed line). Vertical bars represent 95% confidence intervals.

from a trade-off between the risks of starvation at low body mass and predation at high body mass (Covas et al. 2002).

Senescence in Survival of Roe Deer

The best-fitting model included a survival probability constant over time, whereas detection probabilities varied over

time. The naive analysis led to downward-biased, age-specific survival estimates of both sexes when compared with the MR analysis (fig. 2). This systematic bias occurred because deer that are observed for the last time before the end of the study are wrongly considered to be dead by the naive approach, whereas they may actually still be alive but remain undetected (detection probability varied among years, from 0.32 to 0.69; see appendix table A2 for annual estimates of detection probabilities). Senescence in roe deer survival occurred according to both analyses. Most interestingly, we found consistent patterns between sexes in parameter estimates of the Gompertz curve (table 2). The estimated slope b , which is often considered to be a measure of the rate of senescence, was underestimated using the naive analysis by up to 28% in females and 10% in males. In addition, the estimated intercept A , which corresponds to initial mortality, was overestimated by the naive analysis by up to 45% in females and 30% in males.

Discussion

Our study provides clear evidence that flawed evolutionary inferences may occur when the value of individual detectability is <1 (see also Clobert 1995; Martin et al. 1995). We acknowledge that our case studies are based on field data in which the truth is unknown. Although our interpretation of the results seems reasonable, there is no a priori reason that it is actually correct. However, we have shown that detection probabilities in both case studies were <1 and were time varying, and that the incorrect assumption of perfect detection resulted in (i) changing the estimate of the form of natural selection acting on weavers' body mass and (ii) underestimating roe deer rate of survival senescence and overestimating the initial mortality. In agreement with Nichols et al. (1997), we do not claim that the evolutionary studies that formed this assumption are all incorrect, but rather that it is difficult to know whether we should rely on their results if the data have not been reanalyzed using MR models. In addition to assessing the action of natural selection and investigating senescence trends that we have addressed here, MR models can prove to be useful in studying problems such as the fitness associated with alternative reproductive tactics (e.g., Brown and Brown 2004), the evolution of dispersal (e.g., Serrano et al. 2005), or the nature of life-history trade-offs (e.g., Townsend and Anderson 2007).

Nevertheless, although MR models are considered to be as flexible as generalized linear models or survival models that are used when the detection probability value is 1, some specific problems exist in MR models for which particular attention or further research is needed. First, MR models cannot distinguish between survival and emigration (Lebreton et al. 1992), which may lead to biased

Table 2: Parameter estimates (and SEs) of the Gompertz mortality curve of female and male roe deer: naive approach assuming perfect detectability versus mark-recapture (MR) approach

Gompertz parameters	Females		Males	
	Naive analysis	MR analysis	Naive analysis	MR analysis
<i>A</i>	.053 (.008)	.029 (.007)	.123 (.015)	.086 (.016)
<i>b</i>	.155 (.021)	.199 (.027)	.145 (.026)	.160 (.033)

Note: *A*, initial mortality rate; *b*, exponential rate of mortality increase (i.e., rate of senescence).

estimates of selection if emigration is related to phenotypic trait values (e.g., Letcher et al. 2005). Monitoring individuals in multiple sites (so-called multistate MR models; Lebreton and Pradel 2002) can solve this problem by estimating phenotype-dependent movement among sites. Bias may also arise in the calculated strength and form of selection due to covariation between detection probability and the trait under selection (e.g., Janzen et al. 2000). MR models can cope with this issue by relating the detection probabilities to the phenotypic trait values (Kingsolver and Smith 1995).

Second, a challenge lies in possible heterogeneity in survival among individuals. If there are two classes of individuals with low and high survival probabilities, the overall survival rate will be increasingly influenced by the individuals with the highest survival over ages, leading to underestimated or even masked senescence at the population level (Vaupel and Yashin 1985). In the roe deer case study, for example, we removed some individual heterogeneity by analyzing males and females separately, but residual individual variance might still be present. This can be generated by heterogeneity in phenotypic quality (e.g., body mass at independence) but can also be seen as an intrinsic property of a given individual (Link et al. 2002). In standard MR models, assessing the influence of individual covariates on age-specific survival probabilities is possible, but individual random effects are not yet implemented routinely when detectability is <1 (Cam et al. 2002; see, however, Gimenez et al. 2006 and Royle 2008). A new formulation of MR models that disentangle demographic processes (e.g., survival) from their observation (i.e., detection; Pradel 2005; Gimenez et al. 2007) makes us confident that those tools will soon become available.

A further complexity may arise when relationships between traits are neither linear nor quadratic. The form of selection may not be consistent with directional, stabiliz-

ing, or disruptive selection (Conner and Hartl 2004). In the sociable weaver case study, we assumed simple relationships between survival and body mass, although more complex functions might have been more appropriate. When the detection value is assumed to be 1, the use of spline smoothing allows for maximum flexibility in describing the relationship between fitness components and phenotypic traits (Schluter 1988; Schluter and Nychka 1994). Although progress has been made to implement this approach in MR models (Gimenez et al. 2006), new tools are needed to allow for visualization and quantification of natural selection involving multiple traits (O. Gimenez, A. Grégoire, and T. Lenormand, unpublished data).

Overall, we advocate explicit recognition that detection probabilities are often <1 , which can greatly affect conclusions. As a consequence, we believe that the use of an MR modeling framework can be important for proper biological inference. We realize that implementing MR models may be a costly time investment for evolutionary ecologists who have never used these models, but this can be made easier with the use of existing literature (e.g., Lebreton et al. 1992; Williams et al. 2002) and user-friendly software that is available for free (M-SURGE: Choquet et al. 2004; SURPH: Smith et al. 1994; MARK: White and Burnham 1999).

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APPENDIX

The Risk of Flawed Inference in Evolutionary Studies When Detectability Is Less than One

Table A1: Estimated time-dependent detection probabilities of sociable weavers (quadratic model)

Year	Estimated detection	
	probability	SE
1994	.525	.079
1995	.347	.052
1996	.169	.036
1997 ^a	0	...
1998	.124	.045
1999	.829	.085
2000	.790	.075

^a Detection probability in 1997 was fixed to 0 because the birds did not breed at any time in that year and, consequently, no mark-recapture work was performed.

Table A2: Estimated time-dependent detection probabilities of roe deer (Gompertz model)

Year	Females		Males	
	Estimated detection probability	SE	Estimated detection probability	SE
1979	.320	.154	.527	.143
1980	.465	.104	.488	.106
1981	.520	.088	.454	.097
1982	.315	.076	.620	.089
1983	.497	.073	.394	.084
1984	.468	.071	.527	.086
1985	.436	.065	.494	.091
1986	.680	.060	.646	.100
1987	.503	.062	.513	.082
1988	.556	.059	.603	.076
1989	.558	.058	.603	.070
1990	.522	.063	.598	.072
1991	.549	.063	.536	.081
1992	.687	.056	.520	.078
1993	.549	.066	.448	.078
1994	.505	.070	.564	.084
1995	.557	.071	.444	.074
1996	.537	.069	.505	.072
1997	.567	.063	.531	.067
1998	.632	.060	.623	.061
1999	.462	.061	.525	.062
2000	.341	.056	.333	.056
2001	.572	.061	.449	.066
2002	.524	.064	.396	.065
2003	.584	.061	.357	.059
2004	.583	.066	.430	.065
2005	.531	.063	.346	.061

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